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solving for the stationary point $\nabla f(\mathbf{\bar{x}}^*) = 0$ gives the **normal equations**

$$J^{\mathsf{T}}J\mathbf{x}^{\mathbf{\bar{x}}} = -J^{\mathsf{T}}\mathbf{\bar{r}}_{0}.$$

We have three approaches to solving the normal equations for $\bar{\mathbf{x}}^*$ — in increasing order of **computational complexity** and **stability**:

Orthogonal Distance Regression

- (i) Cholesky factorization of $J^T J$,
- (ii) QR-factorization of J, and
- (iii) Singular Value Decomposition of J.

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and $\Phi(\bar{\mathbf{x}}; t_i)$ is our **model**.

The key approximation for the Hessian

Orthogonal Distance Regression

Here, y_i are the **measurements** taken at the **locations/times** t_i ,

 $\nabla^2 \mathbf{f}(\bar{\mathbf{x}}) = J(\bar{\mathbf{x}})^T J(\bar{\mathbf{x}}) + \sum_{i=1}^m r_j(\bar{\mathbf{x}}) \nabla^2 r_j(\bar{\mathbf{x}}) \approx \mathbf{J}(\bar{\mathbf{x}})^T \mathbf{J}(\bar{\mathbf{x}}).$

Summary Orthogonal Distance Regression	Linear Least Squares Nonlinear Least Squares		Summary Orthogonal Distance Regression	Linear Least Squares Nonlinear Least Squares	
Summary: Nonlinear Least Squares		2 of 4	Summary: Nonlinear Least Squares		3 of 4
Line-search algorithm: Gauss-N $\begin{bmatrix} J(\bar{\mathbf{x}}_k)^T J(\bar{\mathbf{x}}_k) \end{bmatrix} \bar{\mathbf{p}}$ Guaranteed descent direction, fast Hessian approximation holds up) of squares problem (used for efficient	lewton , with the subproblem: $\mathbf{b}_{k}^{\text{GN}} = -\nabla f(\mathbf{\bar{x}}_{k}).$ convergence (as long as the equivalence to a linear least \mathbf{z}_{k} , stable solution).		Trust-region algorithm: Leven subproblem: $\bar{\mathbf{p}}_{k}^{\text{LM}} = \arg\min_{\bar{\mathbf{p}} \in \mathbb{R}^{n}} \frac{1}{2} \ J(\bar{\mathbf{x}}_{k})\bar{\mathbf{p}} + \bar{\mathbf{r}}$ Slight advantage over Gauss-New local convergence properties; also least squares problem.	berg-Marquardt , with the $\overline{\sigma}_k \ _2^2$, subject to $\ \overline{\mathbf{p}}\ \le \Delta_k$. where μ to a line \mathbf{p} (locally) equivalent to a line	me ar
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Peter Blomgren, $\langle \texttt{blomgren.peter@gmail.com} \rangle$	Orthogonal Distance Regression	— (5/22)	Peter Blomgren, {blomgren.peter@gmail.com}	Orthogonal Distance Regression	— (6/22)
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Summary: Nonlinear Least Squares

Hybrid Algorithms:

- When implementing Gauss-Newton or Levenberg-Marquardt, we should implement a **safe-guard** for the **large residual case**, where the Hessian approximation fails.
- If, after some reasonable number of iterations, we realize that the residuals are **not** going to zero, then we are better off switching to a general-purpose algorithm for non-linear optimization, such as a quasi-Newton (BFGS), or Newton method.

So far we have assumed that there are **no errors** in the variables describing **where** / **when** the measurements are made, *i.e.* in the data set $\{t_j, y_j\}$ where t_j denote times of measurement, and y_j the measured value, we have assumed that t_j are exact, and the measurement errors are in y_j .

Fixed Regressor Models vs. Errors-In-Variables Models

Under this assumption, the discrepancies between the model and the measured data are $% \left({{{\left[{{{\rm{T}}_{\rm{T}}} \right]}}} \right)$

$$\epsilon_j = y_j - \Phi(\mathbf{\bar{x}}; t_j), \quad i = 1, 2, \dots, m.$$

Next, we will take a look at the situation where we take errors in t_j into account — these models are known as **errors-in-variables models**, and their solutions in the linear case are referred to as **total least squares optimization**, or in the non-linear case as **orthogonal distance regression**.

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ODR = Nonlinear Least Squares, Exploiting Struct

Least Squares vs. Orthogonal Distance Regression



of orthogonal distance regression, where we measure the shortest distance to the model curve. [The right figure is actually not correct, why?]

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Orthogonal Distance Regression: The Weights

The weight-vectors \mathbf{d} and $\mathbf{\bar{w}}$ must either be supplied by the modeler, or estimated in some clever way.

If all the weights are the same $w_i = d_i = C$, then each term in the sum is simply the shortest distance between the point (t_i, y_i) and curve $\Phi(\bar{\mathbf{x}}; t)$ (as illustrated in the previous figure).

In order to get the orthogonal-looking figure, I set $w_i =$ 1/0.5 and $d_i = 1/4$, thus adjusting for the different scales in the t- and y-directions.

The shortest path between the point and the curve will be normal (orthogonal) to the curve at the point of intersection.

We can think of the scaling (weighting) as adjusting for measuring time in fortnights, seconds, milli-seconds, micro-seconds, or SAN DIEGO STAT nano-seconds...

Orthogonal Distance Regression

For the mathematical formulation of orthogonal distance regression we introduce perturbations (errors) δ_i for the variables t_i , in addition to the errors ϵ_i for the y_i 's.

We relate the measurements and the model in the following way

$$\epsilon_j = y_j - \Phi(\mathbf{\bar{x}}; t_j + \delta_j),$$

and define the minimization problem:

$$(\bar{\mathbf{x}}^*, \bar{\delta}^*) = \operatorname*{arg\,min}_{\bar{\mathbf{x}}, \bar{\delta}} \frac{1}{2} \sum_{j=1}^m \bigg[\mathbf{w}_j^2 \bigg[\mathbf{y}_j - \boldsymbol{\Phi}(\bar{\mathbf{x}}; \, \mathbf{t}_j + \delta_j) \bigg]^2 + \mathbf{d}_j^2 \delta_j^2 \bigg],$$

where $\mathbf{\bar{d}}$ and $\mathbf{\bar{w}}$ are two vectors of weights which denote the relative significance of the error terms.

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Orthogonal Distance Regression: In Terms of Residuals r_i

By identifying the 2m residuals

$$r_j(\bar{\mathbf{x}}, \bar{\delta}) = \begin{cases} w_j \left[y_j - \Phi(\bar{\mathbf{x}}; t_j + \delta_j) \right] & j = 1, 2, \dots, m \\ d_{j-m} \delta_{j-m} & j = (m+1), (m+2), \dots, 2m \end{cases}$$

we can rewrite the optimization problem

$$(ar{\mathbf{x}}^*,ar{\delta}^*) = rgmin_{ar{\mathbf{x}},ar{\delta}} rac{1}{2} \sum_{i=1}^m w_j^2 \Big[y_j - \Phi(ar{\mathbf{x}};\ t_j + \delta_j) \Big]^2 + d_j^2 \delta_j^2,$$

in terms of the 2*m*-vector $\mathbf{\bar{r}}(\mathbf{\bar{x}}, \overline{\delta})$

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$$(\bar{\mathbf{x}}^*, \bar{\delta}^*) = \arg\min_{\bar{\mathbf{x}}, \bar{\delta}} \frac{1}{2} \sum_{i=1}^{2m} r_j(\bar{\mathbf{x}}, \bar{\delta})^2 = \arg\min_{\bar{\mathbf{x}}, \bar{\delta}} \frac{1}{2} \|\mathbf{r}(\bar{\mathbf{x}}, \bar{\delta})\|_2^2.$$

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Weighted Least Squares / Orthogonal Distance Regression

Orthogonal Distance Regression \rightarrow Least Squares

If we take a cold hard stare at the expression

$$(\bar{\mathbf{x}}^*, \bar{\delta}^*) = \argmin_{\bar{\mathbf{x}}, \bar{\delta}} \frac{1}{2} \sum_{i=1}^{2m} r_i(\bar{\mathbf{x}}, \bar{\delta})^2 = \arg_{\bar{\mathbf{x}}, \bar{\delta}} \frac{1}{2} \|\bar{\mathbf{r}}(\bar{\mathbf{x}}, \bar{\delta})\|_2^2.$$

We realize that this is now a standard (nonlinear) least squares problem with 2m residuals and (n + m) unknowns — $\{\bar{\mathbf{x}}, \bar{\delta}\}$.

We can use any of the techniques we have previously explored for the solution of the nonlinear least squares problem.

However, a straight-forward implementation of these strategies may prove to be quite expensive, since the number of parameters have doubled to 2m and the number of independent variables have grown from n to (n + m). Recall that usually $m \gg n$, so this is a drastic growth of the problem.



ODR → Least Squares: Exploiting Structure

Fortunately we can save a lot of work by exploiting the structure of the Jacobian of the Least Squares problem originating from the orthogonal distance regression — many entries are zero!

$$\frac{\partial r_j}{\partial \delta_i} = w_j \frac{\partial [y_j - \Phi(\bar{\mathbf{x}}; t_j + \delta_j)]}{\partial \delta_i} = 0, \quad \forall i, j \le m, i \ne j$$

$$\frac{\partial r_j}{\partial \delta_i} = \frac{\partial [d_{j-m}\delta_{j-m}]}{\partial \delta_i} = \begin{cases} 0 & i \ne (j-m), j > m \\ d_{j-m} & i = (j-m), j > m \end{cases}$$

$$\frac{\partial r_j}{\partial x_i} = \frac{\partial [d_{j-m}\delta_{j-m}]}{\partial x_i} = 0, \quad i = 1, 2, \dots, n, \quad j > m$$

Let $v_j = w_j \frac{\partial [y_j - \Phi(\bar{\mathbf{x}}; t_j + \delta_j)]}{\partial \delta_j}$, and let $D = \text{diag}(\bar{\mathbf{d}})$, and $V = \text{diag}(\bar{\mathbf{v}})$, then we can write the Jacobian of the residual function in matrix form...

Orthogonal Distance Regression \rightarrow Least Squares: Problem Size



Figure: We recast ODR as a much larger standard nonlinear least squares problem.

Standard LSQ-solution via QR/SVD ~ $\mathcal{O}(mn^2)$, for $m \gg n$; slows down by a factor of $2(1 + m/n)^2$.

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$ODR \rightarrow Least Squares: The Jacobia$	n

We now have

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$$J(\mathbf{\bar{x}}, \overline{\delta}) = \left[\begin{array}{c|c} \widehat{J} & V \\ \hline 0 & D \end{array}
ight],$$

where D and V are $m \times m$ diagonal matrices, $D = \text{diag}(\mathbf{\bar{d}})$, and $V = \text{diag}(\mathbf{\bar{v}})$, and \widehat{J} is the $m \times n$ matrix defined by

$$\widehat{J} = \left[\frac{\partial [w_j(y_j - \Phi(\bar{\mathbf{x}}; t_j + \delta_j))]}{\partial x_i}\right]_{\substack{j = 1, 2, \dots, m \\ i = 1, 2, \dots, n}}$$

We can now use this matrix in *e.g.* the Levenberg-Marquardt algorithm...

ODR = Nonlinear Least Squares, Exploiting Structur

$ODR \rightarrow Least Squares:$ The Jacobian — Structure



Figure: If we exploit the structure of the Jacobian, the problem is still somewhat tractable.

$ODR \rightarrow Least Squares \rightarrow Levenberg-Marquardt$

If we partition the step vector $\mathbf{\bar{p}}$, and the residual vector $\mathbf{\bar{r}}$ into

$$\mathbf{ar{p}} = \left[egin{array}{c} \mathbf{ar{p}}_x \ \mathbf{ar{p}}_\delta \end{array}
ight], \quad \mathbf{ar{r}} = \left[egin{array}{c} \mathbf{ar{r}}_1 \ \mathbf{ar{r}}_2 \end{array}
ight]$$

where $\mathbf{\bar{p}}_x \in \mathbb{R}^n$, $\mathbf{\bar{p}}_{\delta} \in \mathbb{R}^m$, and $\mathbf{\tilde{r}}_1, \mathbf{\tilde{r}}_2 \in \mathbb{R}^m$, then *e.g.* we can write the Levenberg-Marguardt subproblem in partitioned form

$$\frac{\widehat{J}^{T}\widehat{J} + \lambda I_{n}}{V\widehat{J}} \frac{\widehat{J}^{T}V}{V^{2} + D^{2} + \lambda I_{m}} \left[\frac{\overline{\mathbf{p}}_{X}}{\overline{\mathbf{p}}_{\delta}} \right] = -\left[\frac{\widehat{J}^{T}\widetilde{\mathbf{r}}_{1}}{V\widetilde{\mathbf{r}}_{1} + D\widetilde{\mathbf{r}}_{2}} \right]$$

Since the (2,2)-block $V^2 + D^2 + \lambda I_m$ is diagonal, we can eliminate the $\mathbf{\bar{p}}_{\delta}$ variables from the system...

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$ODR \rightarrow Least Squares \rightarrow Levenberg-Marquardt$ 2 of 2		$ODR \to LSQ\;(2m \times (n+m)) \to Levenberg-Marquardt \to LSQ\;(m \times n)$			

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 $\begin{bmatrix} \frac{\hat{J}^T \hat{J} + \lambda I_n & \hat{J}^T V}{V \hat{J} & V^2 + D^2 + \lambda I_m} \end{bmatrix} \begin{bmatrix} \mathbf{\bar{p}}_x \\ \mathbf{\bar{p}}_\delta \end{bmatrix} = -\begin{bmatrix} \frac{\hat{J}^T \mathbf{\tilde{r}}_1}{V \mathbf{\tilde{r}}_1 + D \mathbf{\tilde{r}}_2} \end{bmatrix}$ $\mathbf{\bar{p}}_{\delta} = -\left[V^2 + D^2 + \lambda I_m\right]^{-1} \left[\left(V\mathbf{\tilde{r}}_1 + D\mathbf{\tilde{r}}_2\right) + V\widehat{J}\mathbf{\bar{p}}_x\right]$

This leads to the $n \times n$ -system $A\mathbf{\bar{p}}_{x} = \mathbf{\bar{b}}$, where

$$A = \left[\widehat{J}^{T}\widehat{J} + \lambda I_{n} - \widehat{J}^{T}V\left[V^{2} + D^{2} + \lambda I_{m}\right]^{-1}V\widehat{J}\right]$$

$$\mathbf{\bar{b}} = \left[-\widehat{J}^{T}\mathbf{\tilde{r}}_{1} + \widehat{J}^{T}V\left[V^{2} + D^{2} + \lambda I_{m}\right]^{-1}\left[V\mathbf{\tilde{r}}_{1} + D\mathbf{\tilde{r}}_{2}\right]\right].$$

Hence, the total cost of finding the LM-step is only marginally more expensive than for the standard least squares problem.

The derived system is typically very ill-conditioned since we have formed a modified version of the normal equations $\widehat{J}^T \widehat{J} + \text{"stuff"}$... With some work we can recast is as an $m \times n$ linear least squares problem $\mathbf{\bar{p}}_{x} = \arg\min_{\mathbf{\bar{p}}} \|\tilde{A}\mathbf{\bar{p}} - \mathbf{\tilde{b}}\|_{2}$, where

$$\tilde{A} = \left[\hat{J} + \lambda [\hat{J}^{T}]^{\dagger} - V \left[V^{2} + D^{2} + \lambda I_{m}\right]^{-1} V \hat{J}\right]$$
$$\tilde{\mathbf{b}} = \left[-\tilde{\mathbf{r}}_{1} + V \left[V^{2} + D^{2} + \lambda I_{m}\right]^{-1} \left[V \tilde{\mathbf{r}}_{1} + D \tilde{\mathbf{r}}_{2}\right]\right]$$

Where the "mystery factor" $[\hat{J}^T]^{\dagger}$ is the **pseudo-inverse** of \hat{J}^T . Expressed in terms of the QR-factorization $QR = \hat{J}$, we have

$$\widehat{J}^T = R^T Q^T, \quad [\widehat{J}^T]^\dagger = Q R^{-T},$$

Since $QR^{-T}R^{T}Q^{T} = I = R^{T}Q^{T}QR^{-T}$.

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	Summary Weighted Least Squares / Orthogonal Distance Regression Orthogonal Distance Regression ODR = Nonlinear Least Squares, Exploiting Structure	Summary Orthogonal Distance Regression	Weighted Least Squares / Orthogonal Distance Regression ODR = Nonlinear Least Squares, Exploiting Structure
oftware and I	References	Index	
MINPACK	Implements the Levenberg-Marquardt algorithm. Available for free from http://www.netlib.org/minpack/.		
ODRPACK	Implements the orthogonal distance regres- sion algorithm. Available for free from http://www.netlib.org/odrpack/.	orthogonal distance regression Jacobian structure, 15 Levenberg-Marquardt formulation, 19	
Other	The NAG (Numerical Algorithms Group) library and HSL (formerly the Harwell Subroutine Library), implement several robust nonlinear least squares implementations.	minimization problem, 10 residuals, 12	
GvL	Golub and van Loan's Matrix Computations, 4th edition (chapters 5–6) has a comprehensive discussion on orthogonalization and least squares; explaining in gory detail much of the linear algebra (<i>e.g.</i> the SVD and QR-factorization) we swept under the rug.		

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Orthogonal Distance Regression

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